

High-Dimensional Covariance Decomposition into Sparse Markov and Independence Domains

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High-Dimensional Covariance Estimation

- n i.i.d. samples, p variables $\mathbf{X} := [X_1, \dots, X_p]^T$.
- High-dimensional regime: both $n, p \rightarrow \infty$ and $n \ll p$.
- Covariance estimation:

$$\Sigma^* := \mathbb{E}[\mathbf{X}\mathbf{X}^T].$$

- Challenge: empirical (sample) covariance ill-posed when $n \ll p$:

$$\widehat{\Sigma}^n := \frac{1}{n} \sum_{k=1}^n \mathbf{x}(k)\mathbf{x}(k)^T.$$

Solution: Imposing Sparsity for Tractable High-dimensional Estimation

Incorporating Sparsity in High Dimensions

Sparse Covariance

$$\begin{bmatrix} & \\ & \end{bmatrix} = \begin{bmatrix} \blacksquare & & & \\ \blacksquare & \blacksquare & & \\ & \blacksquare & \blacksquare & \\ & & & \blacksquare \end{bmatrix}$$

Σ^* Σ_R

Sparse Inverse Covariance

$$\begin{bmatrix} & \\ & \end{bmatrix} = \begin{bmatrix} \blacksquare & & & \\ \blacksquare & \blacksquare & & \\ & \blacksquare & \blacksquare & \\ & & & \blacksquare \end{bmatrix}^{-1}$$

Σ^* J_M^{-1}

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Relationship with Statistical Properties (Gaussian)

- Sparse Covariance= Independence Model: **marginal independence**.
- Sparse Inverse Covariance=Markov Model: **conditional independence**

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Guarantees under Sparsity Constraints in High Dimensions

Consistent Estimation when $n = \Omega(\log p)$ $\Rightarrow n \ll p$.

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Going beyond Sparsity in High Dimensions?

Going Beyond Sparse Models

Motivation

- Sparsity constraints restrictive to have faithful representation.
 - Data not sparse in a single domain
 - **Solution: Sparsity in Multiple Domains.**
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One Possibility: Sparse Markov + Sparse Independence Models

- Sparsity in Multiple Domains: Multiple Statistical Relationships.

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Efficient Decomposition and Estimation in High Dimensions?

Unique Decomposition? Good Sample Requirements?

Summary of Results

$$\Sigma^* = J_M^{*-1} + \Sigma_R^*$$

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Contribution 1: Novel Method for Decomposition

- Decomposition into Markov and residual domains.
 - **Unification** of Sparse Covariance and Inverse Covariance Estimation.
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- Conditions for unique decomposition (**exact statistics**).
- **Sparsistency** and **norm guarantees** in both Markov and independence domains (**sample analysis**)
- **Sample requirement**: no. of samples $n = \Omega(\log p)$ for p variables.

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Efficient Method for Covariance Decomposition and Estimation

Related Works

Sparse Covariance/Inverse Covariance Estimation

- Sparse Covariance Estimation: **Covariance Thresholding**.
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Beyond Sparse Models: Decomposition Issues

- Sparse + Low Rank (Chandrasekaran et. al) (Candes et. al)
 - Decomposable Regularizers (Negahban et. al)
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Our contribution: Guaranteed Decomposition and Estimation

Outline

- 1 Introduction
- 2 Algorithm**
- 3 Guarantees
- 4 Experiments
- 5 Conclusion

Some Intuitions and Ideas

- $\Sigma^* = J_M^{*-1} + \Sigma_R^*$.
- $\hat{\Sigma}^n$: sample covariance using n i.i.d. samples

$$\left[\right] = \left[\begin{array}{c} \text{blue matrix} \\ \text{blue matrix} \\ \text{blue matrix} \end{array} \right]^{-1} + \left[\begin{array}{c} \text{red matrix} \\ \text{red matrix} \\ \text{red matrix} \end{array} \right]$$

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Review Ideas for Special Cases: Sparse Covariance/Inverse Covariance

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Review Ideas for Special Cases: Sparse Covariance/Inverse Covariance

Sparse Covariance Estimation (Independence Model)

- $\Sigma^* = \Sigma_R^*$.

- $\hat{\Sigma}^n$: sample covariance using n samples

$$\left[\begin{array}{c} \\ \\ \end{array} \right] = \left[\begin{array}{c} \\ \\ \end{array} \right]$$

- p variables: $p \gg n$.

- **Thresholding estimator for off-diagonals (Bickel & Levina):** threshold

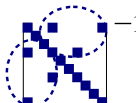
chosen as $\sqrt{\frac{\log p}{n}}$

- **Sparsistency (support recovery) and Norm guarantees** when

$$n = \Omega(\log p) \Rightarrow n \ll p.$$

Recap of Inverse Covariance (Markov) Estimation

- $\Sigma^* = J_M^{*-1}$
- $\hat{\Sigma}^n$: sample covariance using n i.i.d. samples

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$$\hat{J}_M := \operatorname{argmin}_{J_M \succ 0} \langle \hat{\Sigma}^n, J_M \rangle - \log \det J_M + \gamma \|J_M\|_{1,\text{off}}$$

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Max-entropy Formulation (Lagrangian Dual)

$$\hat{\Sigma}_M := \underset{\Sigma_M \succ 0}{\operatorname{argmax}} \log \det \Sigma_M$$

$$\text{s. t. } \|\hat{\Sigma}^n - \Sigma_M\|_{\infty, \text{off}} \leq \gamma, \quad (\Sigma_M)_d = (\hat{\Sigma}^n)_d$$

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Consistent Estimation Under Certain Conditions, $n = \Omega(\log p)$

Extension to Markov+Independence Models?

$$\Sigma^* = J_M^{*-1} + \Sigma_R^*$$

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Sparse Covariance Estimation

Threshold off-diagonal entries of $\hat{\Sigma}^n$.

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Add ℓ_1 **penalty** to maximum likelihood program (involving inverse covariance matrix estimation)

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The diagram illustrates the equation $\left[\begin{array}{c} \\ \end{array} \right] = \left[\begin{array}{c} \\ \end{array} \right]^{-1} + \left[\begin{array}{c} \\ \end{array} \right]$. The first matrix on the right is a blue-bordered matrix with a blue diagonal and blue off-diagonal elements, representing a sparse matrix. The second matrix on the right is a red-bordered matrix with a red diagonal and red off-diagonal elements, representing another sparse matrix. The plus sign is between the two matrices.

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Is it possible to unify above methods and guarantees?

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- Penalties in above methods are in different domains

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Threshold off-diagonal entries of $\hat{\Sigma}^n$.

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- Penalties in above methods are in different domains
- Insight: Consider **dual program of MLE**
- Dual program is in covariance domain for Markov model.

Our Algorithm: Covariance Decomposition

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- Extend ℓ_1 -penalized MLE

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Max-entropy Formulation

- Lagrangian dual of ℓ_1 -penalized MLE

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$$\text{s. t. } \|\hat{\Sigma}^n - \Sigma_M - \Sigma_R\|_{\infty, \text{off}} \leq \gamma, \quad (\Sigma_M)_d = (\hat{\Sigma}^n)_d, \quad (\Sigma_R)_d = 0.$$

$\ell_1 - \ell_\infty$ -penalized MLE (This work)

$$\hat{J}_M := \operatorname{argmin}_{J_M \succ 0} \langle \hat{\Sigma}^n, J_M \rangle - \log \det J_M + \gamma \|J_M\|_{1, \text{off}}$$

$$\text{s. t. } \|J_M\|_{\infty, \text{off}} \leq \lambda,$$

Observations regarding the Proposed Method

$\ell_1 - \ell_\infty$ -penalized MLE (Primal)

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Max-entropy Markov + ℓ_1 -penalized Residuals (Dual)

$$(\hat{\Sigma}_M, \hat{\Sigma}_R) := \operatorname{argmax}_{\Sigma_M \succ 0, \Sigma_R} \log \det \Sigma_M - \lambda \|\Sigma_R\|_{1,\text{off}}$$

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Case: $\lambda \rightarrow 0$ (Sparse Covariance Estimation)

- Threshold estimator for off-diagonals of Σ_R^* (under exact statistics)
- With samples, $\lambda = \sqrt{\log p/n}$ reduces to **threshold estimator**.

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Case: $\lambda \rightarrow \infty$ (Sparse Inverse Covariance Estimator)

- Residual matrix $\hat{\Sigma}_R = 0$: **ℓ_1 -penalized MLE** of Ravikumar et. al

Observations regarding the Proposed Method

$\ell_1 - \ell_\infty$ -penalized MLE (Primal)

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- Residual matrix $\hat{\Sigma}_R = 0$: ℓ_1 -penalized MLE of Ravikumar et. al
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Unification of Sparse Covariance & Inverse Covariance Estimation

Outline

- 1 Introduction
- 2 Algorithm
- 3 Guarantees**
- 4 Experiments
- 5 Conclusion

Guarantees for High-Dimensional Estimation

$$\Sigma^* = J_M^*{}^{-1} + \Sigma_R^*$$

$$\begin{bmatrix} & \\ & \end{bmatrix} = \begin{bmatrix} \text{blue } X & \\ & \text{red } X \end{bmatrix}^{-1} + \begin{bmatrix} \text{red } X & \\ & \text{red } X \end{bmatrix}$$

Conditions for Recovery

- Maximum degree Δ in the Markov graph (corresponding to J_M^*).
- Number of samples n , number of nodes p satisfy $n = \Omega(\Delta^2 \log p)$.
- Regularization constant: $\lambda = \max_{i \neq j} J_M^*(i, j) + \Theta(\sqrt{\log p/n})$.

Theorem

The proposed method outputs estimates $(\hat{J}_M, \hat{\Sigma}_R)$ such that

- $(\hat{J}_M, \hat{\Sigma}_R)$ are **sparsistent** and **sign consistent**.
- satisfy **norm guarantees**.

$$\|\hat{J}_M - J_M^*\|_\infty, \|\hat{\Sigma}_R - \Sigma_R^*\|_\infty = O\left(\sqrt{\log p/n}\right).$$

Observations

Corollary 1 (Sparse Covariance Estimation)

With $\lambda = \Theta(\sqrt{\log p/n})$, our method reduces to threshold estimator (Bickel & Levina) and is sparsistent for covariance estimation.

Corollary 2 (Sparse Inverse Covariance Estimation)

With $\lambda \rightarrow \infty$, our method reduces to ℓ_1 -penalized MLE (Ravikumar et. al) and is sparsistent for inverse covariance estimation.

Conditions for Recovery

- Mutual incoherence-type conditions
- Sample complexity $n = \Omega(\Delta^2 \log p)$.
- Comparable to inverse covariance estimation (Ravikumar et. al).

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Synthetic Data

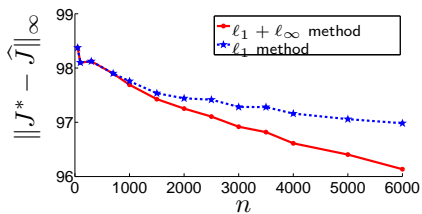
$$\Sigma^* = J_M^{*-1} + \Sigma_R^*, \quad J^* = (\Sigma^*)^{-1}.$$

$$\begin{bmatrix} & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \end{bmatrix}^{-1} = \begin{bmatrix} \text{blue} & & & & & & & \\ & \text{blue} & & & & & & \\ & & \text{blue} & & & & & \\ & & & \text{blue} & & & & \\ & & & & \text{blue} & & & \\ & & & & & \text{blue} & & \\ & & & & & & \text{blue} & \\ & & & & & & & \text{blue} \end{bmatrix} + \begin{bmatrix} & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \\ & & & & & & & \end{bmatrix}$$

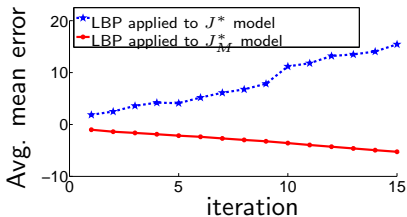
Setup

- 8×8 2-d grid for Markov model.
- Mixed Markov model (both positive and negative correlations).
- Arbitrary-valued sparse residuals.

J estimation



Performance under LBP

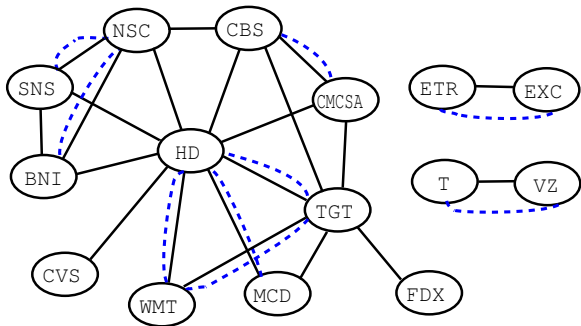


Advantage over existing techniques.

Experiments on Stock Market Data

Setup

- Monthly stock returns of companies on S&P index.
- Companies in divisions **E.Trans**, **Comm**, **Elec&Gas** and **G.Retail Trade**.
- Apply the proposed method.



- Solid line: Markov graph. Dotted line: Independence graph.

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Conclusion

Summary

- Covariance decomposition and estimation in high dimensions
- Combination of Markov and independence models
- Efficient method and guarantees for estimation in both domains
- Unifying sparse covariance/inverse covariance estimation

Conclusion

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Not covered in this talk

- Analysis under **Exact Statistics**: Conditions for **Unique Decomposition**.
- **Sample** Analysis: Careful control of **perturbations in both domains**.

Longer version available on webpage.

Conclusion

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 - Combination of Markov and independence models
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Outlook

- Discrete Model (via pseudo-likelihood)
- Other forms of residuals (e.g. low rank).